

Intriguing yet simple skewness - kurtosis relation in economic and demographic data distributions; pointing to preferential attachment processes

Marcel Ausloos ^{1,2,*} and Roy Cerqueti ³

¹ School of Business, University of Leicester, University Road. Leicester, LE1 7RH, United Kingdom. Email: ma683@le.ac.uk

² GRAPES – Group of Researchers for Applications of Physics in Economy and Sociology. Rue de la Belle Jardinière 483, B-4031, Angleur, Belgium. Email: marcel.ausloos@ulg.ac.be

³ Department of Economics and Law, University of Macerata, Via Crescimbeni 20, I-62100, Macerata, Italy. Tel.: +39 0733 258 3246; Fax: +39 0733 258 3205.

Email: roy.cerqueti@unimc.it.

Abstract

In this paper, we propose that relations between high order moments of data distributions, for example between the skewness (S) and kurtosis (K), allow to point to theoretical models with understandable structural parameters. The illustrative data concerns two cases: (i) the distribution of income taxes and (ii) that of inhabitants, after aggregation over each city in each province of Italy in 2011. Moreover, from the rank-size relationship, for either S or K, in both cases, it is shown that one obtains the parameters of the underlying (hypothetical) modeling distribution: in the present cases, the 2-parameter Beta function, - itself related to the Yule-Simon distribution function, whence suggesting a growth model based on the preferential attachment process.

1 Introduction

Characteristics of distributions of variables is a never ending subject of investigations in many fields of economic research. There is much work for example on testing conditional convergence in variance and skewness for estimating (multivariate or not) normality (Mardia, 1970; Lütkepohl and Theilen, 1991; Nath, 1996; Bai and Ng, 2005; Egger and Pfaffermayr, 2009; Huber and Pfaffermayr, 2010), or on improving goodness-of-fit in regressions (Dufour et al., 2003). In fact, drawing inference on the parameters of (regression or agent based) models is a basic statistical problem which fortunately may provide interesting discoveries (Sahota, 1978; Lin and Lui, 1992; Richardson and Smith, 1993), not only on strict economic problems, but also in related sociological or demographic ones (Eeckhout, 2004; Kaitila, 2014; Vitanov and Ausloos, 2015; Cerqueti and Ausloos, 2015).

There has been already much interesting work on recurrence relations between high moments (μ_i , $i \geq 3$) of order statistics distributions (Arnold et al., 1992; Malik et al., 1988). Knowledge of such moments are of interest for drawing inference about the scaling parameters. Balakrishnan et al. (1988) have reviewed many recurrence relations and identities for several continuous distributions. Pertinently, in view of the following, let us point to Thomas and Samuel (2008) analysis of recurrence relations for the Beta distribution moments, - a distribution of wide application, both in its continuous or discrete form (Johnson et al., 1995; Martinez-Mekler et al., 2009; Ausloos and Cerqueti, 2016a)

The skewness and kurtosis are officially the third and fourth moment, μ_3 and μ_4 respectively, of a distribution, where μ_i is usually centered on the mean; μ_2 is of course the variance, sometimes called σ^2 . Mathematical statistics textbooks and software packages usually calculate the Fisher-Pearson coefficient of skewness and the kurtosis

$$S = \mu_3 / \mu_2^{3/2} \quad (1)$$

$$K = \mu_4 / \mu_2^2. \quad (2)$$

respectively, which by extension become the (commonly accepted measure of) skewness and kurtosis of the distribution, respectively; we use such notations here below.

It is known since Pearson (1916) and Wilkins (1944) that the kurtosis has a theoretical lower bound related to the skewness $K \geq aS^2 + b$. Based on several types of experimental data, one has indeed observed that the quadratic relationship

$$K = p S^2 + q \quad (3)$$

holds as an envelope of scattered data (see references below). In fact, p and q can be empirically fitted constants, which might have some interesting meaning, as pertinently shown by **Guszejnov et al. (2013)**.

It has recently been discussed by **Cristelli et al. (2012)** that a more general form of the $K - S$ relationship could be provided, i.e.

$$K = p S^\nu + q \quad (4)$$

with $\nu = 4/3$. However the findings (and interpretation) have been questioned: such an exponent might be due to the data too limited size (Celikoglu and Tirnakli, 2015)

Much of such experimental data, for which Eq. (3) is obeyed, pertains to magnetohydrodynamics, meteorological and medical data. Moreover, most of the data where such relations are found pertain to time series.

We provide another set of cases in which one finds $\nu = 2$, in an unusual set of data pertinent to a complex geo-sociological-economic realm. Eq. (4), with $\nu \simeq 2$, is found to hold for the wealth and population distribution of Italy cities, aggregated at the provincial level. We tie this specific finding to a statistical process inferring a Beta-distribution, or Polya urn dynamics, as its universal dynamics.

In Section 2, we very briefly recall previous experimental data analysis with similar findings, pertinent to the present report, but obviously in quite (in scientific terms) scattered fields of investigations, - in order to stress the originality of the present ones.

In Section 3, we explain the system complexity which we investigated. Notice at once, that there is no time dependence, therefore the cloud of points, from which other works infer a parabola type relation between S and K , is here strictly "reduced" to a collapsing-like situation, rendering the parameter values being much more precise in view of describing the dynamical process, - if we are allowed to discriminate between (so called) hard and soft science. Such a Section contains also the main methodological investigation methods, results and related comments.

Section 4 is devoted to the description of a theoretical Polya urn model which is related to the developed complex system arguments. Specifically, a preferential attachment system is introduced whose stochastic law follows a Beta distribution, whence whose parameters could be also calibrated for defining skewness and kurtosis of a set of data.

In Section 5, we offer some conclusive remarks and provide also suggestions for future research directions.

Some Tables containing descriptive statistics of the data and the disaggregation of them at a provincial level are reported in Appendix A.

2 Literature review

As mentioned in the introduction, there are several reports pointing to the veracity of Eq. (4), with $\nu = 2$, in various research fields: the greatest occurrence is in turbulence, among the most recent see in magneto-hydrodynamics (Labit et al., 2007; **Guszejnov et al., 2013**; Bergsaker et al., 2015) and in atmospheric physics (see Sura and Sardeshmukh (2008)). In this respect, refer also to Alberghi et al. (2002), where the authors discuss the parameters conditions to be satisfied in the context of air vertical velocity in the atmospheric boundary layer for having a certain relation between skewness and kurtosis. In these cases, in order to have different points in the $S - K$ plane, the authors usually repeat the experiment or evaluate the moments on different time windows of the same series. **For further pertinent references, e.g. see Mole and Clarke (1995) and Sattin et al. (2009).**

Related to the above-mentioned papers, there is a study in the context of geophysics of oil production time series forecasting (Frausto-Solís et al., 2015).

Cristelli et al. (2012) analysed the relation between skewness and kurtosis for earthquakes and daily price returns (on the S&P500) and identified two power-law regimes of non-Gaussianity, on the kurtosis versus skewness plots, but Celikoglu and Tirnakli (2015) demonstrated that the proposed "universal" relation between skewness and kurtosis, in fact is not universal and originates only due to the small number of data points in the data sets considered.

For completeness, on $K - S$ relation consideration in the financial

domain, let us mention related work on $S^2 - K$ bounds for unimodal distributions by Klaassen et al. (2000), and more recently by McDonald et al., (2013) and Kerman and McDonald, (2015) with a discussion about modeling some "popular income distributions" with exponential generalized beta functions.

Other cases where a simple $K - S$ relation is found are in surface roughness analysis (Isoda and Kawashima, 2013; Sharifi-viand et al., 2014). It is also worth mentioning the contributions in medical and biological fields; they pertain to fluctuation responses in the visual cortex (Medina and Diaz, 2016a, 2016b, 2016c) or ventricular fibrillation (Gonzalez-Gonzalez et al., 2013); an apparently comprehensive review about "biological and psychological aspects of the $K-S$ can be found in Cain et al. (2016).

One might put in parallel to the above a paper reporting a $4/3$ relation between skewness and kurtosis of aesthetic score distributions in a photo aesthetics dataset, generated from an online voting survey (Park and Zhang 2015).

Thus, it may be observed that most of the data pertains to time series analysis, many to human reaction time, and a few to various (laboratory or not) produced crystalline samples. However, to the best of our knowledge, not many observations of a peculiar relationship between K and S seem to have been reported on socio-demography aspects.

3 Italy economic and demographic data

A statistical assessment of regional wealth inequalities over Italy (IT) has been previously provided based on aggregated tax income size data (Mir et al., 2014; Cerqueti and Ausloos, 2015a, 2015b, 2015c; Ausloos and Cerqueti, 2016b).

Let it be known that IT is nowadays (since 2010) made up of 8092 cities distributed over 110 provinces. To provide some better understanding of the paper aims and results, the IT administrative structure can be briefly described as follows. Italy is clustered in 20 non-overlapping regions, and each region contains one or more provinces, which in turn are composed by cities along with their territories (*comuni*). Thus, each city belongs to only one province, and each province is contained in only one region.

The Gross Domestic Product (GDP) in Italy was worth about 2276 billion USD in 2011. The population of Italy fell "slightly" below 60 millions, then.

The economic data was obtained from (and by) the Research Center of the Italian Ministry of Economics and Finance (MEF). The population data source is the Italian Institute of Statistics (ISTAT). In particular, data on the population are extracted from the elaborations of the 15th Italian Census, performed by ISTAT in 2011. We have disaggregated contributions at a municipal level for 2011, in order to obtain the aggregated tax income (ATI), ATI_c and the number of inhabitants $N_{inhab,c}$, for each city c .

	$ATI_{c,p}$		$N_{inhab,c,p}$	
	S	K	S	K
Min.	0.58791	-1.3075	0.76111	-0.93186
Max.	17.092	296.36	16.951	292.99
Sum	655.56	4982.1	594.05	4291.0
N_p	110	110	110	110
Mean (μ)	5.9596	45.292	5.4005	39.009
Median (m)	5.7346	35.458	5.0314	27.925
RMS	6.5453	62.251	6.0163	56.184
St. Dev. (σ)	2.7188	42.902	2.6636	40.619
Variance	7.3920	1840.6	7.0948	1649.9
Std Err.	0.25923	4.0905	0.25396	3.8729
Skewn.	0.87472	2.5531	1.0765	2.9083
Kurt.	1.6629	10.238	2.2291	13.121
μ/σ	2.1920	1.0557	2.0275	0.9604
$3(\mu - m)/\sigma$	0.2483	0.6877	0.4157	0.8186
ρ	-0.1234	3.2129	0.0840	4.282
$\mu - 2\sigma$	0.5219	-40.511	0.07328	-42.2298
$\mu + 2\sigma$	11.397	131.09	10.728	120.248

Table 1: Summary of (rounded) statistical characteristics for the distribution of S and K for ATI of cities in IT provinces and for the distribution of the number of inhabitants in IT cities in the various ($N_p = 110$) provinces. **The last two lines will be useful in the assessment of the outliers, see below.**

Equations		$ATI_{c,p}$	$N_{inhab,c,p}$
Eq. (3)	ν (theoretical)	2	2
	p	1.048±0.008	1.052±0.010
	q	0.415±0.499	0.921±0.518
	R^2	0.993	0.991
Eq. (4)	ν (empirical)	1.912±0.026	1.894±0.028
	p	1.324±0.092	1.389±0.103
	q	-2.065±0.900	-1.763±0.898
	R^2	0.994	0.992
Eqs. (11)-(12)	ρ	5.7224	5.9116
	a	0.7556	0.8493
	b	4.9668	5.0623

Table 2: Summary of (rounded) parameters for the quadratic relation between K and S : Eq. (3); for the power law relation between K and S : Eq. (4); for the Beta-distribution model: **Eqs. (11)-(12)** – about ATI of IT cities ($N_c = 8092$) and number of inhabitants in the $N_p = 110$ IT provinces in 2011. For the Beta case, we report a graph of the Beta **cumulative distribution** functions with the relevant calibrated parameters, in Figures 9 and 10. Some related comments can be found in Subsection 4.2. The value of the auxiliary parameter ρ is reported for the sake of completeness. **The \pm refers to the standard error value.**

Eq.(8)	$ATI_{c,p}$		$N_{inhab,c,p}$		Eq.(11)
	K	S	K	S	
κ_4	3.1426±0.525	1.3666±0.075	1.3447±0.244	1.7307±0.993	
γ_4	0.2884±0.009	0.1248±0.005	0.2865±0.009	0.1692±0.005	$b - 1$
ξ_4	0.8853±0.033	0.4816±0.010	1.0434±0.036	0.4378±0.010	$a - 1$
ψ_4	0.2649±0.031	0.1225±0.024	0.1897±0.022	0.2812±0.039	
R^2	0.9947	0.9945	0.9956	0.9951	

Table 3: Summary of (rounded) parameters for the fits to Eq. (8) based on Beta-distribution model about ATI of IT cities in 2011 and number of inhabitants in the $N_p = 110$ IT provinces from the 2007-2011 census; the correspondence between the exponent parameters and the Beta function parameters is given for completeness.

3.1 K - S relation analysis

What we care about is the distributions of the skewness and of the kurtosis: these are the distributions of interest.

The interesting Table is Table 1. It is seen that the corresponding (in some sense "average" S and K are positive and not small. The most relevant point seems to be the existence of (2) negative K values: -1.30750 for BT (Barletta-Andria-Trani, Apulia region) in the ATI case and -0.9319 for BT and -0.7511 for RG (Ragusa, Sicily) for the N_{inhab} .

On the other extreme TO (Torino province in Piedmont region) has the largest K and S for both cases; for information TO contains the largest number of cities (315).

The obtained corner results for these specific provinces are in line with historical and empirical "evidence". In fact, Torino represents the core of the industrial production of Italy, being the headquarter of FIAT. Thus, the related province has an **unequal** distribution of richness and population, with an asymmetry of positive type. It is also expected that the tails of the distribution are heavy, being Torino (the city) one of the largest cities in Italy, – in terms of number of inhabitants and ATI, – and since the same province contains some of the smallest cities of the Country.

For what concerns BT and RG, they are two of the less populated and poor provinces in Italy, with a large part of small cities. Hence, a platycurtic distribution for the ATI and N_{inhab} is what everyone with a fair level of knowledge of the Italian reality should expect.

The K - S relationship for the distribution of ATI of cities in the 110 provinces is given in Fig. 5

The K - S relationship for the distribution of the number of inhabitants in cities in the 110 provinces is shown in Fig. 6. The relation is pretty smooth in both cases, and recall those found in magnetohydrodynamics and other studies of time dependent systems; see pertinent references in Sect.2.

Thereafter, in accord with previous literature, we try two fits: a polynomial of degree 2 but without linear term, i.e. Eq.(3), or the pseudo parabolic polynomial-like form Eq.(4).

The parameters p and q are given in Table 2, together with the corresponding regression coefficient. To leave ν as a free fit parameter is seen not to be a drastic improvement. Thus, one can expect, since $\nu \simeq 2$ that a simple interpretation or modelisation based on well established statistical distribution is in order; see Sect. 4.

It can be noticed that $p \simeq 1$, but q is negative, since there is a negative kurtosis for the distributions in a couple of provinces (see Table 1).

3.2 Rank-size analysis

The above findings remind us that there is a general relationship between skewness and kurtosis within Pearson's distribution system. Therefore, in order to pursue toward some understanding of this sort

of $K - S$ relation, we propose to develop a complementary analysis. Instead of considering the K and S values as belonging to some continuous distribution, we are using a method, the rank-size analysis method, which allows to study a distribution of "quantities" when the orders of magnitudes can be rather different, and when the values have some imprecise error bar, - as it always occurs in such economic and sociological surveys. In such a methodology, the (K and S here) values are supposed to belong to discrete distributions which are regularly sampled. Thus, we write the S and K data in an ascending size (regular) order, independently of each other, for the ATI and the number of inhabitants, respectively, i.e. giving the rank $r = 1$ to the lowest S and to the lowest K values, etc.

The most simple rank-size law is thought to be a power law of the rank r , - leading to the Zipf plot, $y \sim r^{-\alpha}$. It is often modified for including an upper tail cut-off as through the Yule-Simon law,

$$y(r) = d r^{-\alpha} e^{-\lambda r}. \quad (5)$$

In order to take into account a possible change of curvature in the data, if some falling off seems to occur visually, at the highest ranks, Eq. (5) can be then written (Ausloos, 2014a, 2014b; Ausloos and Cerqueti, 2016a) as

$$y_3(r) = \kappa_3 r^{-\gamma} (N - r + 1)^{-\xi}, \quad (6)$$

where $+1$ is introduced in order to avoid a singular point in the fit at the highest rank $r_M = N$, if $\xi \simeq 0$. This also emphasizes that an upper tail toward infinity is rather meaningless, since the upper rank r_M is necessarily finite.

In view of taking into account a better fit at low and high rank, one can further generalize Eq. (6) to a five parameter free equation (Ausloos and Cerqueti, 2016a):

$$y_5(r) = \kappa_5 (r + \Phi)^{-\gamma} (N + 1 - r + \Psi)^{-\xi}, \quad (7)$$

where the parameter Φ is reminiscent of Mandelbrot's generalization of Zipf's law at low rank, while Ψ allows some flexibility at the highest rank, - where usually the error bar on the data can be rather influential in defining r . The shape of the curve in Eq. (7) is sensitive to the variations of Φ and Ψ (Ausloos and Cerqueti, 2016a).

Here, neglecting any low rank free parameter (Φ) of dubious origin, but still allowing for some flexibility on the upper rank value divergence, we approximate Eq. (7) by

$$Y_4(r) = \kappa_4 r^{\xi_4} (N - r + \psi_4)^{-\gamma_4} \quad (8)$$

One is allowed to imagine that a Generalized Discrete Beta function, like Eq.(8), reminds the reader of the Pearson Type I distribution, supported in the relevant rank interval $[0, N]$. In so doing, the corresponding best fits of the various S and K ranked data can be found for the distributions of city *ATI* for 2011 and of the number of inhabitants of the IT

cities according to the 2011 census, distributed over the 110 provinces, respectively.

Fig. 7 displays the rank-size relation for K and S for the distribution of ATI aggregated over cities in the IT 110 provinces in 2011, and the best fits by Eq. (8). In the same spirit, Fig. 8 gives the rank-size relation for K and S , with fits by Eq. (8), for the distribution of the number of inhabitants aggregated over cities in the IT 110 provinces according to the 2011 census. The fit parameters are found in Table 3. The values (and error bars) on the ξ_4 and γ_4 , together with the value of the regression coefficient, R^2 , are quite convincing of the existence of an inflection point in the rank-size data. Nevertheless, observe the variety of values for ξ_4 and γ_4 .

4 Polya Urn Modelization

4.1 Preferential attachment

Our argument for suggesting a model stems from the historical view that cities do not appear nor grow stochastically. Moreover, there is a postulate on demography that ghettos form according to peer status, in particular, due to the wealth of the population: rich and poor group themselves in clusters. In so doing, the number of inhabitants is somewhat related to the wealth. A similar type of process can be imagined, *mutatis mutandis*, for such different qualities: the so called "preferential attachment process". Such a process can be defined as a settlement procedure in urn theory, where additional balls are added and distributed continuously to the urns (cities, in this model) composing the system. The obtained model is the general Polya urn (Mahmoud, 2009). In our context, the rule of such an addition follows an increasing function of the number of balls already contained in the urns. The settlement formation obeys a Yule process, with a log-normal initial distribution of the population of the settlements.

In general, such a process contemplates also the creation of new urns. In such a general framework, this model is associated to the Yule-Simon distribution (Vitanov and Ausloos, 2012, 2015) whose density function f is

$$f(a; b) = b B(a, b + 1), \quad (9)$$

where a is a positive integer, $b > 0$, $B(a; b)$ is the Euler Beta function

$$B(a; b) = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}, \quad (10)$$

$\Gamma(x)$ being the standard Gamma function (Abramowitz and Stegun, 1970; Gradshteyn and Ryzhik, 2000). Explicitly,

$$B(a, b) = \int_0^1 x^{a-1}(1-x)^{b-1} dx \quad (11)$$

denotes the *Beta*-function; a random variable X is Beta-distributed if its probability density function (pdf) obeys

$$f(x) = \frac{x^{a-1} (1-x)^{b-1}}{B(a,b)}. \quad (12)$$

In practical words, newly created urn starts out with k_0 balls and further balls are added to urns at a rate proportional to the number k that they already have plus a constant $a \geq -k_0$. With these definitions, the fraction $P(k)$ of urns (areas) having k balls (cities) in the limit of long time is given by

$$P(k) = \frac{B(k+a; b)}{B(k_0+a; b-1)} \quad (13)$$

for $k \geq 0$ (and zero otherwise). In such a limit, the preferential attachment process generates a long-tailed distribution following a hyperbolic (Pareto) distribution, i.e. a power law, in its tail.

4.2 K and S parametrization

The relevant Beta-function moments, i.e., K and S , are given by Johnson and Kotz (1970, pages 40-44), and recalled by Hanson (1991), in terms of a and b parameters of the Beta function for the normalized variables:

$$K = 2(b-a) \frac{\sqrt{a+b+1}}{\sqrt{ab}(a+b+2)} \quad (14)$$

$$S = 3(a+b+1) \frac{2(a+b)^2 + ab(a+b-6)}{ab(a+b+2)(a+b+3)}; \quad (15)$$

In order to develop the algebra, one also introduces a so called "help variable" (Hanson, 1991)

$$\rho = 6 \frac{K - S^2 - 1}{6 + 3S^2 - 2K} \quad (16)$$

Notice that if Eq. (3) holds, then

$$\rho = 6 \frac{(p-1)S^2 + (q-1)}{(3-2p)S^2 + 2(3-q)}, \quad (17)$$

allowing a theoretical estimate at once, e.g. if p and q are as in Table 2, and a possible comparison to empirical results, shown on the last line in Table 5. In fact, $\rho = a + b$. Thus, one can obtain a and b from:

$$ab = a(\rho - a) = \frac{6\rho^2(\rho + 1)}{(r\rho + 2)(\rho + 3)K - 3(\rho - 6)(\rho + 1)} \quad (18)$$

which leads to 2 solutions for a :

$$a = \frac{\rho}{2} \left[1 \pm \sqrt{1 - \frac{24(\rho + 1)}{(\rho + 2)(\rho + 3)K - 3(\rho - 6)(\rho + 1)}} \right] \quad (19)$$

If the skewness is positive then the larger solution will be the value of b otherwise the larger solution will be the value of a (**Hanson, 1991**). The relevant values of the best fitted Beta can be read from Table 3.

Figures 9 and 10 illustrate the **cumulative distribution** functions of the calibrated Beta distribution for the normalized ATI and N_{inhab} cases, respectively. We notice that the shapes of the distributions are similar, and suggest a high concentration around the small values of the Beta distribution.

5 Conclusion

This paper explores the relationship between skewness S and kurtosis K for series of demographical and economical data. The considered sample is taken from the Italian National Institute of Statistics (number of inhabitants) and the Ministry of Economics and Finance (income taxes) for the Italian cities, aggregated at a provincial level; the reference year is 2011. The existence of a quadratic and a power law relation between skewness and kurtosis is illustrated, with fits which are both visually quite appealing and markedly statistically sound. These findings support and add to the empirical literature on the connections between K and S ; this is the first time that such quadratic and power laws rules are found for socio-economic surveys.

It should seem interesting to search for the general conditions leading to distributions with such apparently simple relations.

In our presently investigated "socio-economic case", the empirical results have been supported by a theoretical argument based on the Polya urn. In so doing, one gets the data-driven calibration of the parameters of a Beta distribution, which adds further insights on the different nature of economic and demographic data.

Appendix A: Descriptive statistics

Table 4 lists the ATI (in EUR) of each provinces in IT, given in alphabetical order of their legal acronym, the province population, and for general information, the number of cities in the relevant province, in 2011.

In Table 5, one displays a summary of the (rounded) statistical characteristics for the distribution of city ATI $ATI_{c,p}$ of each IT provinces ($N_p = 110$) in 2011, for the distribution of the number of inhabitants in a given province, $N_{inhab,c}$ and for information the number of cities in the relevant province, $N_{c,p}$.

In Table 5, we also give σ/μ , called the coefficient of variation (CV), allowing to have some confidence in a relatively peaked distribution, - if CV is not too

large. For completeness, we also provide the immediately deduce value of an indirect measure, $3(\mu - m)/\sigma$.

Figures 1-4 provide a view of the empirical distributions of the skewness and kurtosis for either the ATI or the number of inhabitants of the IT provinces. The shapes of the distributions for S and for K are quite similar. In both cases some outliers emerge, **according to the definition of outliers as those values outside the interval $(\mu - 2\sigma, \mu + 2\sigma)$; see the values reported in Table 1 and just compare them with the histograms in Figures 1-4.**

Notice that both distributions of S and K seem to be more concentrated around their mean values for N_{inhab} rather than for ATI , hence suggesting a more evident regularity in the asymmetry and in the peaks of the distribution across provinces in the former case. **Moreover, Figures 1-4 show also that skewness and kurtosis of our specific dataset, unexpectedly, seem to depart from a normal asymptotic distribution.**

prov	<i>ATI</i>	Ninhab	$N_{c,p}$	prov	<i>ATI</i>	Ninhab	$N_{c,p}$	prov	<i>ATI</i>	Ninhab	$N_{c,p}$
AG	2.844722	447310	43	FR	4.488987	493928	91	PU	4.235173	363003	6
AL	5.738100	428417	190	GE	13.45363	862267	67	PV	7.602668	537620	19
AN	6.176045	475038	49	GO	1.973022	139983	25	PZ	3.137922	378409	10
AO	1.873498	126982	74	GR	2.651100	221442	28	RA	5.322050	384575	1
AP	2.286054	209887	33	IM	2.422774	212854	67	RC	4.151080	547897	9
AQ	3.149092	297418	108	IS	0.808795	87578	52	RE	7.297694	518011	4
AR	4.227090	344453	39	KR	1.030377	171331	27	RG	2.210240	308329	1
AT	2.709870	217870	118	LC	5.033768	336705	90	RI	1.681281	156142	7
AV	3.343223	430292	119	LE	6.149541	803554	97	RM	59.68562	4042676	19
BA	11.31215	1248086	41	LI	4.414274	336412	20	RN	3.812157	322294	2
BG	14.54681	1087401	244	LO	3.102357	224393	61	RO	2.871922	242409	5
BI	2.540291	182417	82	LT	5.154685	544391	33	SA	8.297128	1091227	19
BL	2.798485	210277	69	LU	4.860075	388922	35	SI	3.740127	267194	3
BN	2.134726	285677	78	MB	12.98964	841102	55	SO	2.211207	181091	7
BO	16.17945	981807	60	MC	3.615307	319181	57	SP	2.968193	220063	3
BR	3.225014	401207	20	ME	5.437395	653470	108	SR	3.196919	397952	2
BS	15.63700	1240553	206	MI	55.71135	3072152	134	SS	3.303827	329616	6
BT	2.536561	391127	10	MN	5.256847	408893	70	SV	3.797833	282255	6
BZ	7.469868	505067	116	MO	9.965410	687237	47	TA	5.131752	579836	2
CA	5.858126	552303	71	MS	2.388878	200387	17	TE	2.907707	305872	4
CB	0.200671	226982	84	MT	1.637990	200842	31	TN	7.262721	526510	2
CE	6.270454	906600	104	NA	23.17133	3058592	92	TO	32.29885	2245252	3
CH	3.758159	388280	104	NO	5.194036	364217	88	TP	3.187134	430843	2
CI	1.111378	128581	23	NU	1.318428	158456	52	TR	2.742176	228944	3
CL	1.871670	273155	22	OG	0.457826	57492	23	TS	3.709890	233077	0
CN	7.447040	586599	250	OR	1.331566	164113	88	TV	11.37026	877905	9
CO	8.109499	587547	160	OT	1.496835	151627	26	UD	7.305619	536035	19
CR	4.857371	357473	115	PA	10.25093	1239837	82	VA	12.30706	873241	14
CS	5.017154	715485	155	PC	4.073974	284711	48	VB	1.953740	160385	7
CT	8.251406	1080034	58	PD	12.33167	921659	104	VC	2.381702	176853	8
CZ	2.839421	360165	80	PE	3.319864	315629	46	VE	11.29438	850523	4
EN	1.157228	173668	20	PG	7.750671	657535	59	VI	11.06162	859987	19
FC	5.071875	390381	30	PI	5.434720	412729	39	VR	11.77778	903564	9
FE	4.787859	353725	26	PN	4.150770	310983	51	VS	0.342572	101396	2
FG	4.562346	627007	61	PO	3.101017	246219	7	VT	3.264761	313998	6
FI	14.30128	971437	44	PR	6.737250	428652	47	VV	1.119749	161952	5
FM	1.793660	175047	40	PT	3.368717	288415	22				

Table 4: Summary table of the ATI (in EUR), number of inhabitants and number of cities of provinces in IT (in e+09 units) in the reference year 2011. Acronyms for provinces are those legally used, in IT.

$ATI_{c,p}$		$N_{inhab,c,p}$		$N_{c,p}$	
min.(x10 ⁻⁵)	3.3479	min.(x10 ⁻⁴)	5.7492	Min.	6
Max.(x10 ⁻¹⁰)	4.5490	Max.(x10 ⁻⁶)	4.0276	Max.	315
Sum(x10 ⁻¹¹)	7.2184	Sum(x10 ⁻⁷)	5.9571	Sum	8092
mean(μ)(x10 ⁻⁷)	8.9204	mean(μ)(x10 ⁻⁵)	5.4155	Mean	73.564
median(m)(x10 ⁻⁷)	2.4601	median(m)(x10 ⁻⁵)	3.7131	Median	60
RMS(x10 ⁻⁸)	6.7701	RMS(x10 ⁻⁵)	7.9654	RMS	91.902
Std.Dev.(σ)(x10 ⁻⁸)	6.7115	Std.Dev.(σ)(x10 ⁻⁵)	5.8680	Std.Dev.	55.338
Var.(x10 ⁻¹⁷)	4.5044	Var.(x10 ⁻¹¹)	3.4433	Var.	3062.27
Std.Err.(x10 ⁻⁶)	7.4609	Std.Err.(x10 ⁻⁶)	0.5595	Std.Err.	5.2762
Skewn.	49.490	Skewn.	3.6571	Skewn.	1.7294
Kurt.	2994.7	Kurt.	15.873	Kurt.	3.6845
μ/σ	0.1329	μ/σ	0.9229	μ/σ	1.3294
$3(\mu - m)/\sigma$	0.2889	$3(\mu - m)/\sigma$	0.8703	$3(\mu - m)/\sigma$	0.7353
r	2.4020	r	0.6254	r	-0.2417

Table 5: Summary of (rounded) statistical characteristics for the distribution of ATI of cities in IT provinces, for the distribution of inhabitants in IT cities in the various ($N_p = 110$) provinces, and for the distribution of the number of cities in provinces in 2011.

Figure 1: Histogram of the skewness of the ATI's for the reference year 2011 and for the IT provinces.

Figure 2: Histogram of the kurtosis of the ATI's for the reference year 2011 and for the IT provinces.

Figure 3: Histogram of the skewness of the number of inhabitants for the reference year 2011 and for the IT provinces.

Figure 4: Histogram of the kurtosis of the number of inhabitants for the reference year 2011 and for the IT provinces.

Figure 5: K - S relationship, Eq. (4), for the distribution of ATI aggregated over cities in the IT 110 provinces in 2011; $\nu \simeq 1.91$.

Figure 6: K - S relationship, Eq. (4), for the distribution of the number of inhabitants of the IT cities in the 110 provinces according to the 2011 Census; $\nu \simeq 1.89$.

Figure 7: Rank-size relation for K and S , with fits with Eq. (8), for the distribution of ATI aggregated over cities in the IT 110 provinces in 2011; fit parameters in Table 3.

Figure 8: Rank-size relation for K and S , with Eq. (8) best fits, for the distribution of the number of inhabitants of the IT cities in the 110 provinces according to the 2011 census; fit parameters in Table 3.

Figure 9: **Cumulative distribution** function of the Beta distribution, with parameters a and b calibrated according to the ATI data for the IT provinces, in the reference year 2011, **and empirical distribution of the data**. Fit parameters are in Table 2.

Figure 10: **Cumulative distribution** function of the Beta distribution, with parameters a and b calibrated according to the demographic data *Ninhab* for the IT provinces, in the reference year 2011, **and empirical distribution of the data**. Fit parameters are in Table 2.

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